



MANICALAND STATE UNIVERSITY
OF
APPLIED SCIENCES

FACULTY OF APPLIED SCIENCES & TECHNOLOGY

DEPARTMENT OF APPLIED STATISTICS

MODULE: PROBABILITY THEORY II

CODE: ASTA 124

SESSIONAL EXAMINATIONS

OCTOBER 2023

EXAMINER: MRS S MANDIZVIDZA

INSTRUCTIONS

1. Answer **All** in Section A.
2. Answer **three** questions in Section B.
3. Start a new question on a fresh page.
4. Total marks: 100.

Additional material(s)

- Statistical tables, Non-programmable electronic scientific calculator, List of formulae.

SECTION A [40 MARKS]

Answer **ALL** questions in this section

A 1

Given the Poisson random variable (Y/λ) ,

$$f(x) = \begin{cases} \frac{e^{-\lambda} \lambda^y}{y!} & \text{for } y = 0, 1, 2, 3, \dots \\ 0 & \text{for } otherwise \end{cases}$$

(a) Show that the moment generating function (mgf), (3)

$$M_Y(t) = e^{\lambda(e^t - 1)}$$

(b) Hence find the (3)

1. first moment $E(Y)$.
2. second moment $E(Y^2)$.

(c) Hence show that mean of $Y = \lambda$ and $var(Y) = \lambda$. (2)

A 2

Let $X = (X_1, \dots, X_n)^T$ be a random vector and $t = (t_1, t_2, \dots, t_m)^T \in \mathfrak{R}^n$.

(a) Define the moment generating function for a random vector X . (2)

Now assume that $X = (X_1, \dots, X_n)^T$ are independent random vectors in \mathfrak{R}^n and let $X = (X_1, \dots, X_n)^T$.

(b) Prove that the moment generating function of the random vector X (defined as a sum) is given by $M_X(t) = \prod_{i=1}^m M_{X_i}(t)$. Assume that X is a random vector in \mathfrak{R}^n , A is an $m \times n$ real matrix and $b \in \mathfrak{R}^m$. (5)

(c) Prove that the moment generating function of a related random vector, $Y = AX + b$ is given at $t \in \mathfrak{R}^m$ by $M_Y(t) = e^{t^T b} M_X(A^T t)$ (5)

A 3

(a) Given that a random variable X has a characteristic function $\phi_X(t)$, Prove that the characteristic function of a related random variable $Y = aX + b$ is given by $\phi_Y(t) = e^{ibt} \phi_X(at)$ where i is a complex number such that $i^2 = -1$. (8)

- (b) If X and Y are independent random variables with characteristic functions $\phi_X(t), \phi_Y(t)$ respectively, prove that the characteristic function of a random variable $Z = X + Y$ is given by $\phi_Z(t) = \phi_X(t)\phi_Y(t)$ (6)

A 4

The probability density function (p.d.f) of a random variable X which has a Cauchy distribution is given by

$$f(x) = \begin{cases} \frac{1}{\pi(1+x)} & \text{for } -\infty < x < 0 \\ 0 & \text{for } otherwise \end{cases}$$

- (a) Show that $E(X)$ does not exist. (6)

SECTION B [60 MARKS]

Answer any **THREE** questions in this section

A 5

Suppose that Y follows a standard normal distribution with mean 0 and variance 1 and probability density function, pdf given by

$f_Y(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}$ for $-\infty < y < \infty$ Now consider the transformation $U = g(Y) = Y^2$. Find the probability density of U using each of the specified methods.

- (a) The transformation (Jacobian) transformation technique. (10)
- (b) The Cumulative distribution Function, CDF technique. (10)
- (c) The moment generating function, m.g.f technique. (7)
- (d) Identify the resulting distribution(s). (3)

A 6

Let the pdf of X_1 and X_2 be given by $f_{X_1, X_2}(x_1, x_2) = 2 \exp[-(x_1 + x_2)]$ Consider two random variables Y_1 and Y_2 be defined in the following manner. $Y_1 = 2X_1, Y_2 = X_2 - X_1$

- (a) Find the joint density of Y_1 and Y_2 and also prove that Y_1 and Y_2 are independent. (15)

Given that X and Y are independent random variables such that $X \sim \exp(\frac{1}{2})$, that is each has probability density function $\frac{1}{2} \exp^{-\frac{x}{2}}$ for $x > 0$.

- (b) find the probability density of $\frac{X+Y}{2}$. (7)
- (c) Suppose that $f(x) = \frac{x}{2}$ for $0 \leq x < 2$ and $g(y) = 2(1 - y)$ for $0 \leq y < 1$. Determine the function $y(x)$ which will transform $f(x)$ into $g(y)$ (8)

A 7

Let Z_1, \dots, Z_n be independent random vectors such that $Z_i \sim N(0, 1)$.

- (a) Prove that the multivariate moment generating function, of $Z = (Z_1, \dots, Z_n)^T$ is given as $M_Z(t) = e^{(t^T \frac{t}{2})}$ (5)
- (b) Hence derive the moment generating function of $X = AZ + \mu$, where A is an $n * n$ real matrix and $\mu \in \mathbb{R}^n$. If $X = (X_1, \dots, X_n)^T \sim N(\mu, \Sigma)$ where Σ is a non-singular definite defined as $\Sigma = A * A^T$ with joint probability density function $f_X(x) = \frac{1}{\sqrt{(2\pi)^n \det \Sigma}} \exp^{-\frac{1}{2}(X - \mu)^T \Sigma^{-1}(X - \mu)}$ (6)

In the case for the bivariate case(i.e for $n = 2$), $\mu = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}$

$$\Sigma = \begin{bmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{bmatrix} \text{ where } \mu_i = E(X_i) \text{ and } \text{Var}(X_i) = \sigma_i^2 \text{ for } i = 1, 2$$

(c) Prove that the joint probability density of $(X_1, X_2)^T \sim N(\mu, \Sigma)$ is given by (9)

$$f_{X_1, X_2}(x_1, x_2) = \frac{1}{(2\pi\rho\sigma_1\sigma_2\sqrt{1-\rho^2})} * \exp\left[-\frac{1}{2(1-\rho^2)}\right] \left[\left(\frac{X_1-\mu_1}{\sigma_1}\right)^2 + \frac{(X_2-\mu_2)^2}{(\sigma_2^2)} - \frac{2\rho(X_1-\mu_1)(X_2-\mu_2)}{(\sigma_1\sigma_2)} \right]$$

Suppose that the time (in minutes) required to serve a customer at a certain store has an exponential distribution with a mean of 3.

(d) What is the probability that the time to serve a customer will exceed 3.75 minutes? (5)

(e) Use the Central Limit Theorem to determine the approximate probability that the total time to serve a random sample of 16 customers will exceed 1 hour (5)

A 8

Consider two random variables Y_1 and Y_2 be defined in the following manner. $Y_1 = 2X_1$
 $Y_2 = X_2 - X_1$

(a) Find the joint density of Y_1 and Y_2 and also prove that Y_1 and Y_2 are independent. (10)

(b) find the probability density of $\frac{X-Y}{2}$. (7)

Suppose that $f(x) = \frac{x}{2}$ for $0 \leq x < 2$ and $g(y) = 2(1 - y)$ for $0 \leq y < 1$.

(c) Determine the function $y(x)$ which will transform $f(x)$ into $g(y)$ (8)